

Semiparametric Analysis of German East–West Migration – Facts and Theory

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Introduction

- German East-West migration in 1991
- microdata from the German Socio Economic Panel (**GSOEP**), $n = 3367$
- Generalized Linear Model (**GLM**) does not fit the data
- semiparametric Generalized Partial Linear Model (**GPLM**) reveals nonlinear influence of household income on migration propensity
- this nonlinear influence is compatible with the **option value** approach of Burda (95)



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Classical Economic Theory

- income is key determinant of migration
- difference between income the host region W^W and income at home W^E at time t (1991: $t = 0$): $\Omega_t = W_t^W - W_t^E$
- forward-looking agent will consider expected net present value (**ENPV**) =
 - EPV of income from migrating
 - EPV of income from not-migrating
 - fixed costs of migrating
- under standard assumptions ENPV is a linear function of current (=1991) income differential Ω_0 .



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Example:

Ω_t follows Brownian motion with drift ν :

$$d\Omega_t = \nu dt + \sigma dz_t$$

where $dz_t = \epsilon_t \sqrt{dt}$, $\epsilon_t \sim N(0, 1)$.

$$\Rightarrow \text{ENPV} = V^m = \frac{1}{\delta} (\Omega_0 + \nu/\delta) - F$$

where δ is the rate of discount.

Marshallian decision rule

$$Y = 1 \quad \text{if } V^m > 0$$
$$Y = 0 \quad \text{otherwise}$$

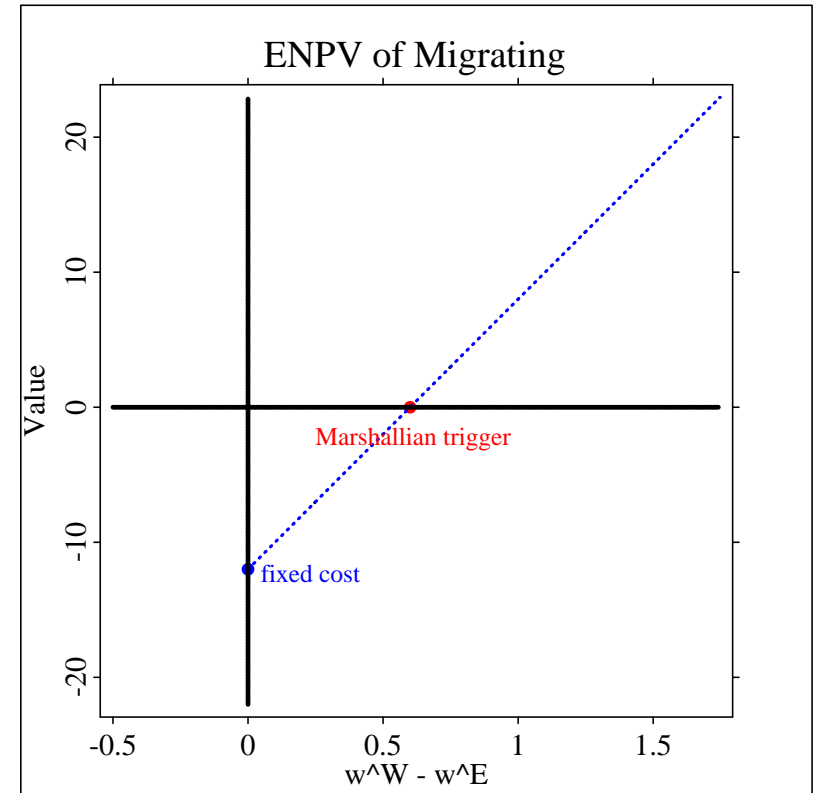


Figure 1: Marshallian theory of migration



The Data

- 3367 observations from GSOEP's 2nd East-German wave (spring of 1991)
- dependent variable Y : *migration propensity*
- measuring current income differential Ω_0 :
imputation is prone to error (self-selection, unemployed, out of the labor force)
include income in East (W_0^E) only
- 11 explanatory variables
- All calculations were done in [XploRe](http://www.xploRe-stat.de)®.
See <http://www.xploRe-stat.de>



Summary statistics

		Mean	Expected Effect
Y	migration intention	.39	
X_1	female	.51	
X_2	partner	.85	–
X_3	owner	.32	–
X_4	family/friends in west	.85	–
X_5	unemp./jobloss certain	.20	+
X_6	env. satisfaction	3.9	–
X_7	city size < 10,000	.52	
X_8	city size 10-10,000	.34	
X_9	university degree	.08	
X_{10}	age min: 18, max: 65	39.4	–
X_{11}	household income min: 200, max: 4000	2189.5	



Parametric Estimation Results

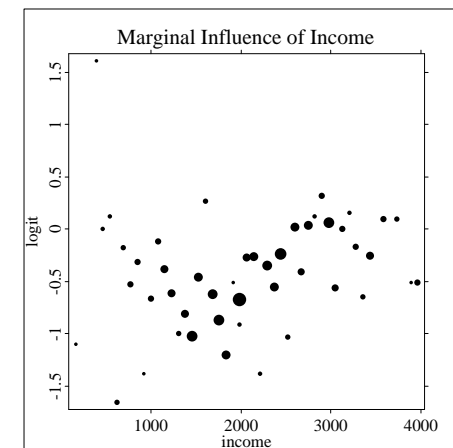
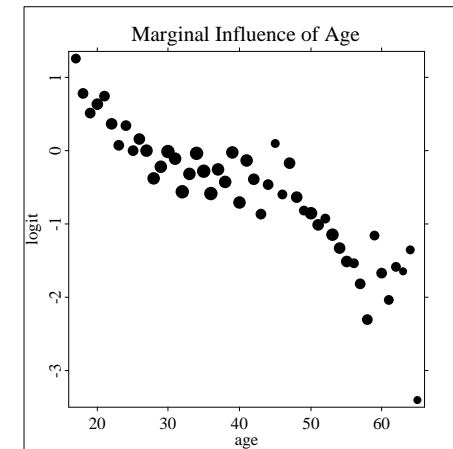
GLM (Logit) estimates of β in $E[Y|x] = 1/\{1 + \exp(-\beta^T x)\}$.

dependent variable: migration intention		
Variable	$\hat{\beta}$	t
constant	1.864	7.74
female	-.233	-3.03
partner	-.325	-2.87
owner	-.576	-5.79
family/friends in west	.647	5.61
unemployed	.217	2.24
env. satisfaction	-.057	-3.52
city size < 10,000	-.718	-5.69
city size 10-100,000	-.347	-2.91
university degree	.481	3.56
age	-.050	-14.89
household income	.0001202	2.22
sample size: 3367, log likelihood: -1992.7		



income & age: linear or nonlinear ?

age and income vs. the logits $\log\{\hat{p}/(1 - \hat{p})\}$



Semiparametric Model

Latent-variable assumption

$$\text{GLM: } Y = 1 \quad \text{if } Y^* = x^T \beta + \alpha t + \alpha_0 - u > 0$$

$$\text{GPLM: } Y = 1 \quad \text{if } Y^* = x^T \beta + m(t) - u > 0$$

t : income in the East (W_0^E)

Distributional assumption

$$\text{GLM \& GPLM: } F_{u|x,t}(\bullet) = \frac{1}{1 + \exp(-\bullet)}$$

GPLM:

$$E(Y|x, t) = \frac{1}{1 + \exp[-\{x^T \beta + m(t)\}]}$$



Semiparametric Estimation

- $\hat{\beta}$ can be found for known m ,
- \hat{m} can be found for known β .

Iterative algorithm (Link function !) employs:

- "usual" likelihood for β

$$\mathcal{L}(\beta) = \sum_{i=1}^n L \{x_i^T \beta + m_\beta(t_i); y_i\}$$

- "smoothed" likelihood for $m(t)$

$$\mathcal{L}^S \{m_\beta(t)\} = \sum_{i=1}^n K_h(t - t_i) L \{x_i^T \beta + m_\beta(t); y_i\}$$

Severini & Staniswalis (1994), Severini & Wong (1992),
Hastie & Tibshirani (1990), Speckman (1988)



Semiparametric Estimation Results

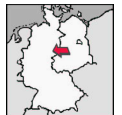
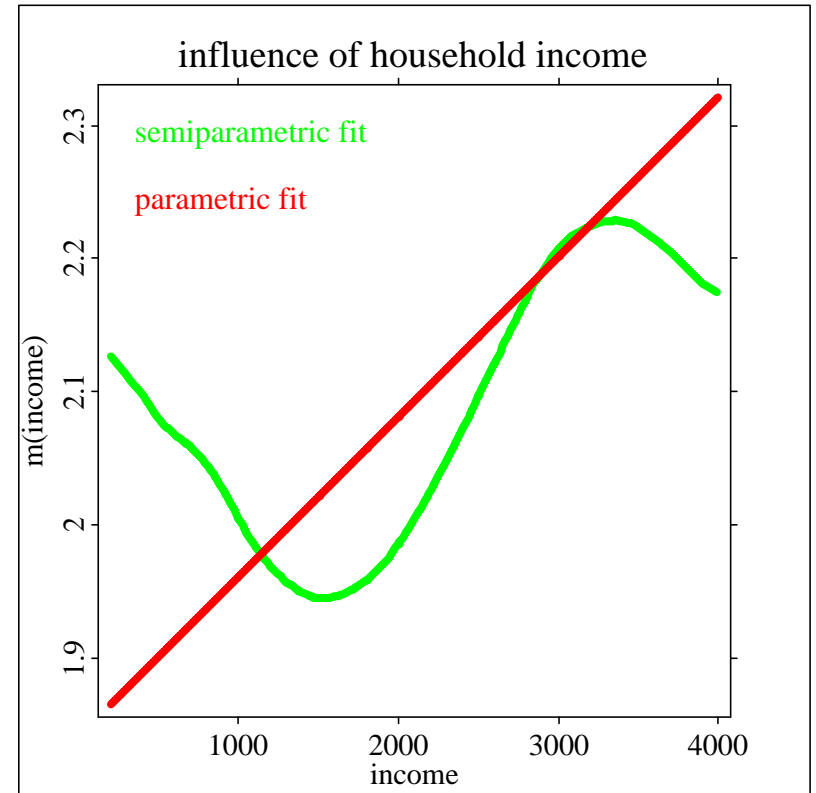
dependent variable: migration intention				
Variable	GPLM		Logit	
	$\hat{\beta}$	t	$\hat{\beta}$	t
female	-.238	-3.1	-.233	-3.0
partner	-.282	-2.4	-.325	-2.9
owner	-.569	-5.7	-.576	-5.8
family/friends in west	.640	5.5	.647	5.6
unemployed	.216	2.2	.217	2.2
env. satisfaction	.056	-3.5	-.057	-3.5
city size < 10,000	-.689	-5.4	-.718	-5.7
city size 10-10,000	-.323	-2.7	-.347	-2.9
university degree	.471	3.5	.481	3.6
age	-.050	-14.9	-.050	-14.9

sample size: 3367, log likelihood: -1989.8, $h = 0.3$

GPLM estimates are close to Logit counterparts



estimated influence of income: $\hat{m}(t)$



Semiparametric Specification Testing

test that $m(t)$ is a linear function:

$$H_0 : m(t) = \alpha t + \alpha_0,$$

$$H_1 : m(t) \text{ is an arbitrary smooth function,}$$

Likelihood ratio test (Hastie & Tibshirani, 1990)

$$R = 2 \sum_{i=1}^n \{ L(\hat{\mu}_i, y_i) - L(\tilde{\mu}_i, y_i) \}$$

$$\text{semiparametric: } \hat{\mu}_i = G\{x_i^T \hat{\beta} + \hat{m}(t_i)\}$$

$$\text{parametric: } \tilde{\mu}_i = G\{x_i^T \tilde{\beta} + \tilde{\alpha} t + \tilde{\alpha}_0\}$$



\hat{m} has a non-negligible smoothing bias



Modified likelihood ratio test

bias-corrected parametric estimate

$$\bar{m}(t_j)$$

from

$$\{G(x_i^T \tilde{\beta} + \tilde{\alpha} t_i + \tilde{\alpha}_0), x_i, t_i\}, \quad i = 1, \dots, n$$

modified LR statistic

$$R^M = 2 \sum_{i=1}^n \{ L(\hat{\mu}_i, \hat{\mu}_i) - L(\bar{\mu}_i, \hat{\mu}_i) \}$$

$$\text{where } \bar{\mu}_i = G\{x_i^T \tilde{\beta} + \bar{m}(t_i)\}$$

Härdle, Mammen & Müller (1996)



asymptotically equivalent

$$\tilde{R}^M = \sum_{i=1}^n w_i \left\{ x_i^T (\hat{\beta} - \tilde{\beta}) + \hat{m}(t_i) - \bar{m}(t_i) \right\}^2$$

with

$$w_i = \frac{[G' \{x_i^T \hat{\beta} + \hat{m}(t_i)\}]^2}{V[G \{x_i^T \hat{\beta} + \hat{m}(t_i)\}]}$$

Asymptotic Normality

Under linearity hypothesis

$$(i) \quad R^M = \tilde{R}^M + o_p(v_n),$$

$$(ii) \quad v_n^{-1}(R^M - e_n) \xrightarrow{D} N(0, 1),$$

where

$$e_n = \left\{ \lambda_T \cdot \int K(u)^2 du \right\} \{h_1 \dots h_q\}^{-1},$$

$$v_n^2 = 2 \left[\lambda_T \int \{K \star K(u)\}^2 du \right] \{h_1 \dots h_q\}^{-1},$$



Bootstrap works

It holds

$$d_K(R^{M*}, R^M) \xrightarrow{P} 0$$

where d_K denotes the Kolmogorov distance.

1. Generate samples y_1^*, \dots, y_n^* with

$$E^*(y_i^*) = G(x_i^T \tilde{\beta} + \tilde{\alpha} t_i + \alpha_0)$$

2. Calculate estimates based on the bootstrap samples and finally the test statistics R^{M*} . The quantiles of the distribution of R^M are estimated by the quantiles of the conditional distributions of R^{M*} .



Test Results

h	0.1	0.2	0.25	0.3	0.4
R	0.028	0.021	0.019	0.017	0.016
R^M	0.053	0.069	0.130	0.269	0.602
R^{M*}	0.015	0.005	0.005	0.005	0.010

- clear rejection of the linearity hypothesis across all bandwidths for R and the bootstrapped R^{M*} .
- The normal approximation for R^M works bad for higher bandwidth levels (Müller, 1997)

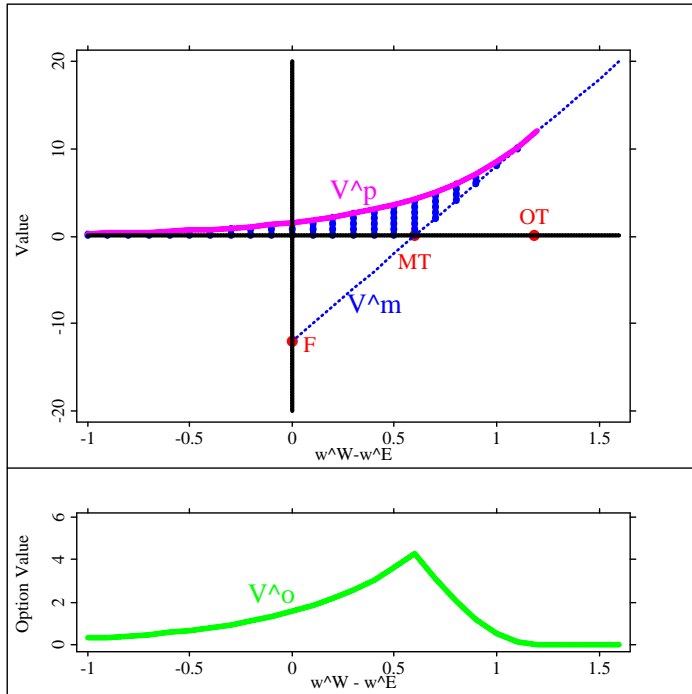


Theoretical Explanation I

Option Value of the Migration Investment

- Marshallian theory: migration occurs now or never.
- Dixit and Pindyck (1994): postponement of the decision without forsaking it can be a valuable option
- delaying migration: more information can be acquired while fixed cost can be avoided
- migrating today means forgoing the opportunity to postpone migration
option value of waiting V^o .
- V^o = what one is willing to pay for the option to postpone the migration decision rather than having to decide now or never





V^p : expected net present value from postponing migration

V^m : expected net present value from migrating today

V^o : option value of waiting

Marshallian decision rule

$$Y = 1 \quad \text{if } \frac{1}{\delta} (\Omega_0 + \nu/\delta) - F > 0$$

$$Y = 0 \quad \text{otherwise}$$

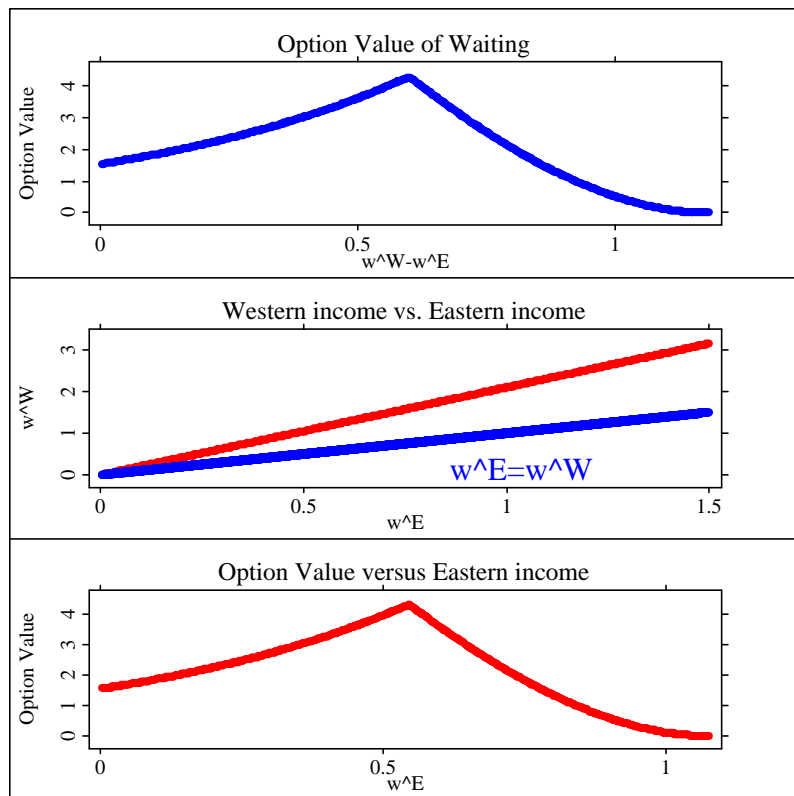
Option value decision rule

$$Y = 1 \quad \text{if } \frac{1}{\delta} (\Omega_0 + \nu/\delta) - F - V^o(\Omega_0) > 0$$

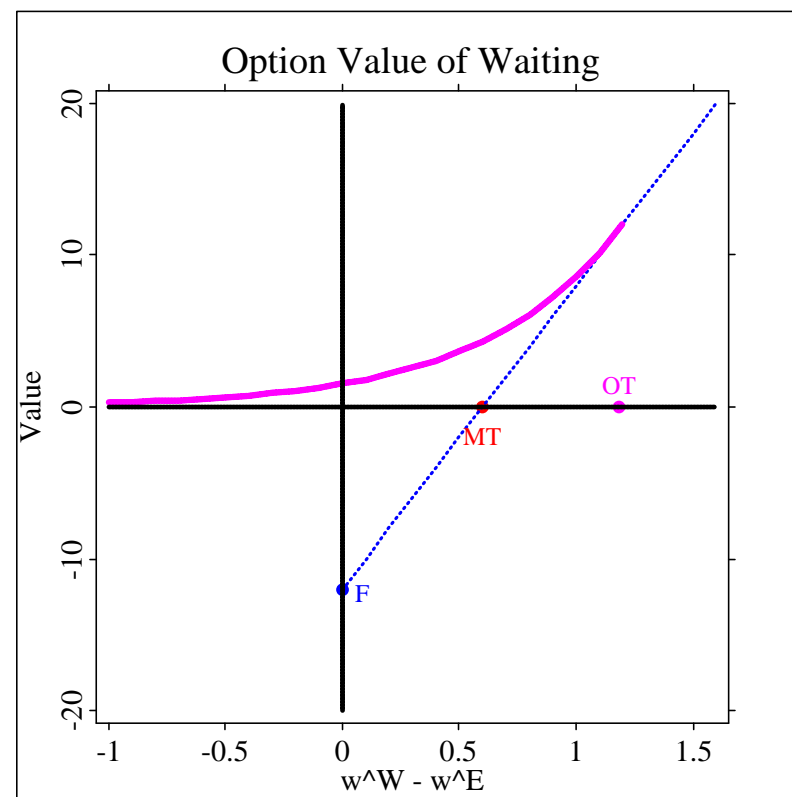
$$Y = 0 \quad \text{otherwise}$$



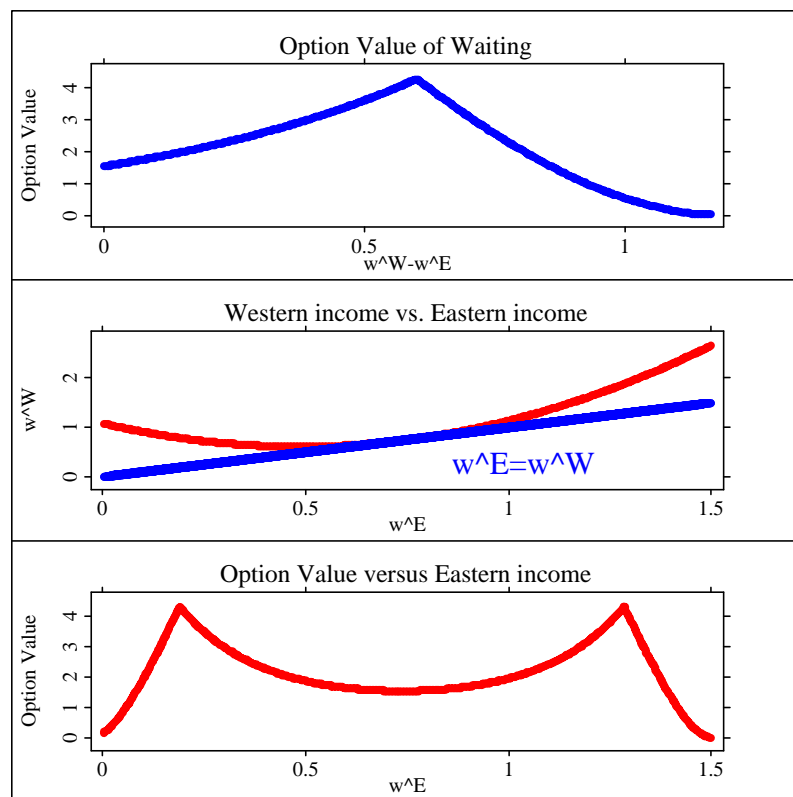
income *differential* versus income in East.



Theoretical Explanation II



income *differential* versus income in East.



Conclusions

- empirical analysis of the propensity to migrate using microdata from the **GSOEP**
- parametric **GLM** did not fit the data
- semiparametric **GPLM** fit produced U-shaped relation between income and migration propensity
- **U-shaped** relation **significantly** deviates from linearity
- estimated influence may be explained by a number of alternative determinants of migration, including the recently proposed **option-value-of-waiting** theory

